

Effect of the choice of food composition table on nutrient estimates: a comparison between the British and American (Chilean) tables

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Submitted 30 April 2003; Accepted 5 August 2003

Abstract

Objective: To determine the level of agreement between the American (Chilean) and British food composition tables in estimating intakes of macronutrients and antioxidants.

Design, setting and subjects: Information based on a food-frequency questionnaire with emphasis on antioxidants was collected from 95 Chileans aged 24–28 years. Nutritional composition was analysed using the British table of food composition and the American table of food composition modified by Chilean food items. Mean differences and limits of agreement (LOAs) of estimated intake were assessed.

Results: Mean differences between the two tables of food composition ranged from 5.3% to 8.9% higher estimates when using the American (Chilean) table for macronutrients. For micronutrients, a bias towards a higher mean was observed for vitamin E, iron and magnesium when the American (Chilean) table was used, but the opposite was observed for vitamin A and selenium. The intra-class correlation coefficient (ICC) ranged from 0.86 (95% confidence interval (CI) 0.81–0.91) to 0.998 (95% CI 0.995–1.00), indicating high to excellent agreement. LOAs for macronutrients and vitamins A and C were satisfactory, as they were sufficiently narrow. There was more uncertainty for other micronutrients.

Conclusion: The American table gives relative overestimates of macronutrients in comparison to the British table, but the relative biases for micronutrients are inconsistent. Estimates of agreement between the two food composition tables provide reassurance that results are interchangeable for the majority of nutrients.

Keywords
Food composition tables
Correlation
Agreement
Macronutrients
Micronutrients

The quality of food consumption and food composition data must be considered simultaneously for the analysis and interpretation of food intake in a population^{1,2}. The quality of food composition data is greatly influenced by the control of variation in food composition, the accuracy in the description of nutrients, the methods of analysis and the mode of data expression³.

Problems with these elements make the food composition table a known major source of error in the estimation of nutrient intakes⁴. Differences and errors are also found when different nutrient databases based on the same national food composition table are compared⁵. International comparisons are more complex, as other differences are added. The most common are related either to the foods (different systems to name, group and describe them) or to the analysis (insufficient description of foods, missing values, ambiguous nomenclature and different chemical analytical methods for specific nutrients)^{6,7}. These lead not only to food composition

databases that differ substantially in form, content and relative quality, but also to bias in the interpretation of diet–disease relationships⁷.

A large amount of work has been done to produce more detailed and homogeneous food composition tables, which has provided some guidelines on the organisation and content of nutritional databases, methods for compiling tables and procedures for the accurate international exchange of data^{8,9}. However, many countries, both developed and developing, still lack appropriate national food composition tables^{6,10–12}. This has motivated the use of tables from other countries with more accurate or complete information, either as the local database reference for nutritional data analysis¹¹ or for the substitution of missing values for specific nutrients¹².

Some advantages may be extracted from using other tables, such as an improvement in the quality of nutritional analysis and less expense and loss of time than in the production of new tables¹². However, comparability is a

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major issue, as local conditions involved in the production of foods will certainly affect their chemical composition¹³. Furthermore, the identification of foods might be particularly difficult for countries that are very rich in 'typical' foods or meals, and where imputed values taken from similar foods may be of uneven reliability^{12–14}.

Many researchers in the field have direct experience with only one set of tables of food composition. There is little knowledge about the extent to which the results of an analysis may be coloured by the peculiarities of a particular set of tables. This paper addresses the question of whether the choice of a particular set of tables may have an impact on the results by assessing the levels of agreement and comparability of estimates using two nutritional databases: the first based mainly on the American table of food composition and including some Chilean nutritional data, the other based on the British table of food composition.

Subjects and methods

Population

The study was carried out in a sub-sample of 95 people, as part of a cohort study that aimed to explore the relationship between diet and asthma in a semi-urban area of Central Chile. Adults between 24 and 28 years old were selected to complete a food-frequency questionnaire (FFQ) with special emphasis on the consumption of antioxidants and fatty acids to determine the dietary intakes of these foods.

Food-frequency questionnaire

The quantitative FFQ was designed to obtain information on the consumption of antioxidants and fatty products during the past month. It included a variety of foods that cover around 90% of the common food intake in the Chilean population¹⁵.

The FFQ was divided into 12 food groups, containing a total of 65 items: fruits with high levels of vitamin C and flavonoids; vegetables rich in vitamins A, C and E and flavonoids; legumes; animal meat (chicken, beef, poultry, pork); fish with high levels of fat (salmon, tuna) and shellfish; eggs; cereals (bread, pasta, rice, sweet biscuits, cakes); fatty products (offal, margarine, butter); dairy foods (whole or skimmed milk, cheese, yoghurt, milky desserts); sugar, ham, honey, sweets; red wine; and non-alcoholic drinks such as juices, tea and coffee. An additional question was added on the consumption of nutritional supplements.

For each group of food, the frequency of consumption was recorded as daily, weekly, fortnightly or monthly. Food intake was first recorded in the FFQ as 'standard portion sizes' (teaspoons, cups, units, etc.). These portions were then translated into grams for each food item. After calculating the total grams consumed daily, the data were analysed using each food composition database.

Food composition tables (databases)

To obtain estimates using the American (Chilean) database, data on food consumed were analysed with a computer program based on information provided by the US Department of Agriculture's (USDA) Nutrient Database, which includes information on energy and 28 nutritional components of more than 5000 foods¹⁶. The surveys that contributed to this database were the Nationwide Food Consumption Survey and the Continuing Survey of Food Intakes by Individuals conducted by the USDA, and the National Health and Nutrition Examination Survey conducted by the Department of Health and Human Services. This source of nutritional composition has been the most common nutritional reference for dietary surveys in Chile^{17,18}.

For some Chilean staple foods that do not appear in the American database, and for specific supplementations, chemical nutritional analysis provided by the eighth edition of the *Chilean Table of Foods Chemical Composition*¹⁹ (1992 version) was added. It contains a description of macro- and micronutrients for a wide range of foods produced or harvested in the country and for traditional Chilean meals.

For the purposes of this study, the Chilean table was considered for the nutritional composition of the following foods:

- Bread, which alone represents on average almost 40% of the energy intake of the Chilean population¹⁷. Wheat flour has a particular composition, since in 1965 a national compulsory fortification system including vitamin A, niacin and riboflavin was approved¹⁵.
- Updated content of omega 6 ($n-6$) and omega 3 ($n-3$) fatty acids (poly- and monounsaturated)²⁰.

The estimates using the British table of food composition were obtained using the program IDA (Integrated Dietary Analysis; IDA Ltd, London, 1997), which is based on the Royal Society of Chemistry's Database 1988–1995 plus all available supplements and appendices²¹. Many of the values included in the British tables were derived from a series of analytical studies commissioned by the Ministry of Agriculture, Fisheries and Food. These studies were done by purchasing samples of foods in local or regional shops or other retail outlets, or incorporated other validated information from the literature on foods of similar characteristics to those found in Britain.

Estimates of nutrients using each of the two sets of food composition tables were independently obtained, in Chile for the USA (Chile) table of food composition and in the UK for the British table of food composition. Although the British tables are not adapted to the peculiarities of the Chilean diet this did not constitute a major problem, as only one item of food, *longaniza*, was unavailable and finally entered as 'sausage'.

Recommendations

For the nutritional recommendations the reference values used were those given by the British Department of Health²², based on the World Health Organization recommendations²³, which apply to the Chilean population included in this survey. The Estimated Average Requirement (EAR) is defined as the notional mean requirement, assuming that the distribution of requirements for a nutrient in a group of individuals is symmetrical.

Statistical analysis

Mean intakes of energy and nutrients estimated using the American (Chilean) nutritional database were compared with those obtained from the British using the paired *t*-test.

Agreement between estimates was expressed by limits of agreement (LOAs), as recommended by Bland and Altman²⁴. The Bland–Altman method provides an absolute measure of agreement, in units of each measurement. This does not allow easy comparison across nutrients that have different units of measurement and variation in intake. The intra-class correlation coefficient (ICC) is a measure of relative repeatability²⁵. If the two estimates from the American (Chilean) table and the British table measure the same quantity then they should be replicated, so the ICC was calculated to measure

relative agreement. It is suggested that an ICC of 0.6 is the lowest limit of a useful measurement²⁶.

Pearson's correlation coefficient was calculated to assess the linear relationship between the two estimates for each individual of the same measurement reported using the two databases. It is an inappropriate measure of agreement²⁴, but many researchers have used it for this purpose and we provide it here to show the difference.

Results

Nutritional intakes of total energy, macro- and micronutrients, as estimated using each database, are presented in Table 1. The protein and micronutrient intakes of participants were well above the EAR values, while the consumption of saturated fatty acids was below the EAR. The medians estimated from the American (Chilean) table were consistently higher for total energy and macronutrients than those from the British table. The differences were particularly large for carbohydrates and protein. In the subgroups of fatty acids, only the intake of polyunsaturated fatty acids (PUFA) was higher when using the American (Chilean) table.

Differences in median values of micronutrients based on estimates from the two tables of food composition were less consistent (Table 1). Medians for vitamin E, iron and magnesium were higher when estimates were based on

Table 1 Energy and nutrient intakes according to the American (Chilean) and British tables of food composition. Values are expressed as median (interquartile range)

Nutrient	EAR*	American (Chilean)	British
Energy (kcal day ⁻¹)	2245	2358 (1941–3434)	2250 (1785–3325)
Carbohydrates			
%TEI	60	58.7	57.7
g day ⁻¹	337.2	351.6 (288.8–513.5)	320.0 (250.1–482.2)
Protein			
%TEI	10		
g day ⁻¹	50.2	101.2 (73.1–142.7)	89.0 (64.2–135.1)
Total fat			
%TEI	30	27.9	28.4
g day ⁻¹	75	73.4 (52.7–112.5)	72.1 (50.6–103.4)
Saturated fatty acids			
%TEI	12	7.5	8.0
g day ⁻¹	35	18.1 (12.7–28.8)	19.4 (13.1–28.0)
Monounsaturated fatty acids			
%TEI	10	10.3	11.1
g day ⁻¹	25	25.5 (17.6–42.6)	27.4 (18.4–41.2)
Polyunsaturated fatty acids			
%TEI	6	7.6	7.0
g day ⁻¹	15	20.9 (15.5–31.1)	17.3 (12.5–25.3)
Vitamin A (µg day ⁻¹)	650	1503 (940–2422)	1597 (1034–2465)
Vitamin C (mg day ⁻¹)	25	158.8 (105.1–280.7)	152.0 (103.0–266.0)
Vitamin E (mg day ⁻¹)	5	16.1 (13.1–22.8)	15.6 (12.4–22.0)
Iron (mg day ⁻¹)	10.1	18.1 (14.0–25.3)	14.6 (11.9–21.3)
Magnesium (mg day ⁻¹)	225	305.5 (216.0–377.4)	280.9 (189.7–349.6)
Selenium (µg day ⁻¹)	ND	98.4 (72.4–143.2)	106.2 (77.9–156.0)
Zinc (mg day ⁻¹)	6.4	8.6 (6.0–10.7)	9.0 (6.3–11.1)

EAR – Estimated Average Requirement; %TEI – percentage of total energy intake; ND – not defined.

* The values given are means of the requirements for men and women.

the American (Chilean) table, but vitamin A and selenium were higher when using the British table.

Table 2 shows the mean differences and LOAs for energy and nutrients. The mean differences and *P*-values between the estimates from the two tables of food composition confirm the results described in Table 1. The bias towards higher values from the American food composition table, expressed as a percentage of the overall mean for each nutrient, varied from 5.2% for total fats to 8.7% for carbohydrates, but for PUFA was 15.6%. The bias was also high for iron (19.0%). LOAs are given for each nutrient. The LOAs are sufficiently narrow for most macronutrients, except PUFA. They were also narrow for vitamins A and C. There is more uncertainty in the interpretation of the LOAs for iron, magnesium, selenium, zinc and vitamin E. Figure 1 gives examples of Bland–Altman plots. The LOAs are narrow for the three nutrients, except that for selenium the discrepancy between food tables depended on the level of intake.

ICCs ranged between 0.86 for iron and 0.99 for vitamins A and C (Table 3). Although the ICCs are all very high they are highest for all macronutrients, except PUFA, and for vitamins A and C, confirming our impression from the LOA results. Pearson's correlation coefficient is also given; it ranged from 0.91 (zinc) to 0.99 (e.g. vitamin A) (Table 3), but is higher than the ICC and gives a misleading representation of agreement, particularly when relative bias was detected.

Discussion

Main findings

Higher estimates for macronutrients were observed when the American (Chilean) table was used. Mean differences for micronutrients were biased, but the direction of the bias was inconsistent. The ICC for each nutrient was very

high, ranging from 0.86 to 0.99, indicating – contrary to expectations – that the level of agreement between the two tables of food composition is excellent for most nutrients or at least satisfactory. The ICC estimates the between-individuals variance in relation to the total variance including error, and thus shows that cross-sectional analysis of macro- and micronutrients in relation to diet should be relatively robust to the choice of database.

Limitations

Our study was based on an FFQ designed to estimate dietary intakes of specific antioxidants and only a limited number of foods were included. However, this FFQ gathers information on the most usual foods, and it represents more than 90% of the daily energy intake of the studied population¹⁵.

Comparison with other studies

Food composition tables that have their own specific conversion factors for the determination of energy^{27,28} and use different methods for analysing carbohydrates¹⁶ may be a source of variation in the estimation of nutrients. In the case of the American and British tables, values for energy and macronutrients are expressed using the same units or at least units that are convertible (joules and calories, pounds and grams)^{7,17,27}. However, in the USDA food composition tables, carbohydrates are calculated including fibre, and digestibility is a factor taken into account when calculating conversion factors into energy¹⁶. Instead, in the British table, available carbohydrates are expressed as monosaccharides and do not consider digestibility when converting into energy^{7,21}, which could explain the different values obtained for energy and carbohydrates in our study. Other international comparisons have reported that nutritional

Table 2 Level of agreement for daily intakes of energy, macro- and micronutrients, according to Chilean (American) and British tables of food composition

Nutrient	Mean difference (95% CI)	<i>P</i> -value	Mean*	% bias†	LOA (reference range for difference)
Energy (kcal)	182.1 (161.2 to 203.0)	<0.001	2610.5	7.0	–23.0 to 387.3
Carbohydrates (g)	33.1 (28.4 to 37.8)	<0.001	380.2	8.7	–12.9 to 79.1
Protein (g)	9.5 (8.3 to 10.7)	<0.001	106.2	8.5	–2.5 to 21.6
Total fat (g)	4.3 (3.3 to 5.3)	<0.001	81.8	5.2	–5.5 to 14.2
Saturated fatty acids (g)	0.14 (–0.14 to 0.7)	0.69	27.7	0.5	–5.3 to 5.6
Monounsaturated fatty acids (g)	–0.1 (–0.7 to 0.4)	0.33	31.1	–0.3	–0.6 to 5.7
Polyunsaturated fatty acids (g)	3.2 (2.7 to 3.7)	<0.001	20.5	15.6	–1.8 to 8.3
Vitamin A (mg)	–72.3 (–81.2 to –63.3)	<0.001	1899.9	–3.8	–160.1 to 15.4
Vitamin C (mg)	2.2 (–1.0 to 5.4)	0.90	228.5	1.0	–29.5 to 33.9
Vitamin E (mg)	0.8 (0.4 to 1.2)	<0.001	17.5	4.6	–3.2 to 4.9
Iron (mg)	3.5 (3.0 to 3.9)	<0.001	18.4	19.0	–1.1 to 8.1
Magnesium (mg)	20.6 (12.0 to 29.2)	<0.001	304.5	6.8	–64.0 to 105.3
Selenium (µg)	–9.5 (–12.1 to –6.9)	<0.001	115.2	–8.3	–35.0 to 15.9
Zinc (mg)	–0.1 (–0.4 to 0.2)	0.24	9.0	–1.1	–3.1 to 2.9

CI – confidence interval; LOA – limit of agreement.

* Mean = mean of two means estimated from the two food composition tables.

† % bias = mean difference divided by the mean.

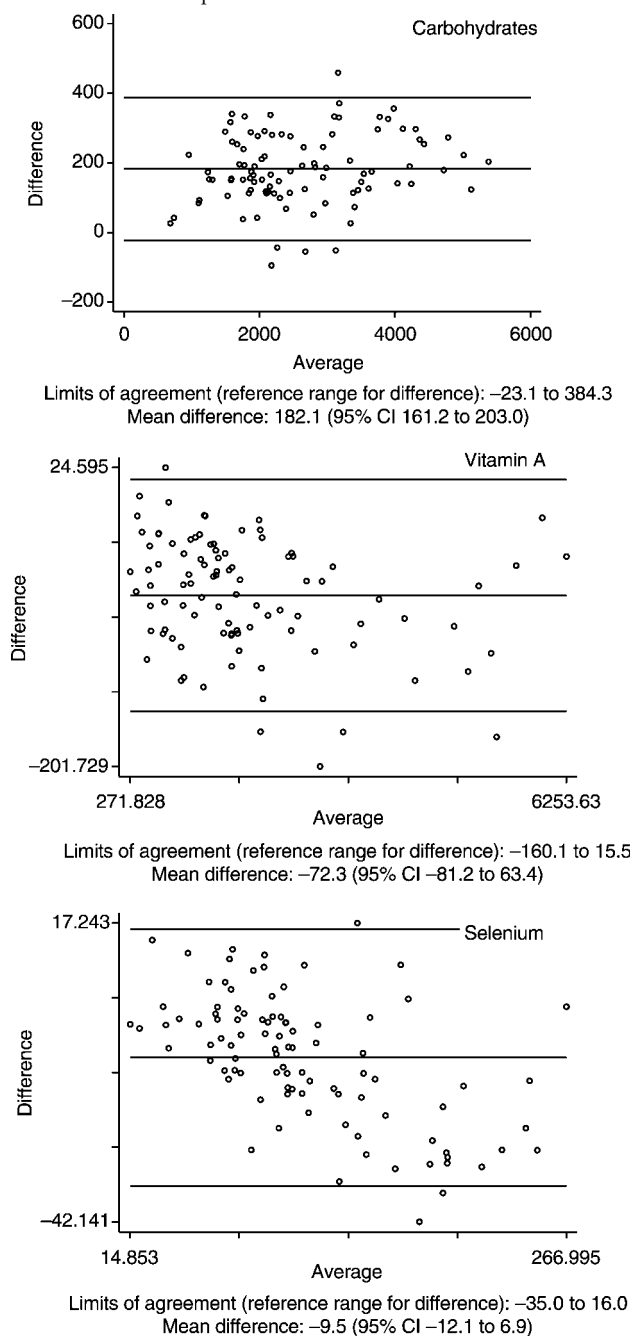


Fig. 1 Bland–Altman plots of mean differences and limits of agreement for carbohydrates, vitamin A and selenium (the difference is given as the value using the American (Chilean) table of food composition minus the value using the British table of food composition). CI – confidence interval

databases greatly influenced by the USDA data, like Central and South American tables^{7,29}, have given values of carbohydrates up to 29% higher than those estimated using European databases, when comparing the same foods.

Problems related to the differences between methods for calculating energy and nutrients have been suggested as reasons for lack of comparability between food composition databases around the world. Although these

Table 3 Intra-class correlation coefficient and Pearson's correlation coefficient for energy, macro- and micronutrients

Nutrient	Intra-class correlation coefficient (95% CI)	Pearson's correlation coefficient
Energy	0.98 (0.97–0.98)	0.99
Protein	0.96 (0.95–0.98)	0.99
Carbohydrates	0.97 (0.95–0.98)	0.98
Total fat	0.98 (0.98–0.99)	0.99
Saturated fatty acids	0.98 (0.97–0.99)	0.98
Monounsaturated fatty acids	0.98 (0.98–0.99)	0.98
Polyunsaturated fatty acids	0.92 (0.89–0.93)	0.97
Vitamin A	0.998 (0.995–1.00)	0.99
Vitamin C	0.995 (0.992–0.998)	0.99
Vitamin E	0.95 (0.94–0.97)	0.96
Iron	0.86 (0.81–0.91)	0.96
Magnesium	0.94 (0.92–0.96)	0.95
Selenium	0.95 (0.94–0.97)	0.98
Zinc	0.91 (0.88–0.94)	0.91

CI – confidence interval.

issues have been discussed since the 1940s³⁰, they still remain as unsolved issues when comparing nutritional databases³. National food composition tables are primarily aimed at providing nutrient data over time at a local level, so that they are not necessarily conceived to provide internationally comparable data. This has been shown in a large research study on food composition tables from nine European countries, where problems in common methods and definitions for some nutrients (folate, dietary fibre) or modes of expression (energy, protein, carbohydrates, carotenes, vitamins A and E) have not yet been resolved⁶. Given the number of reasons that can explain differences between tables of food composition, it is pleasing to note that our study, while confirming the bias between estimates from the British and American tables, shows a high level of agreement between the two. Thus the choice of food composition table, American or British, may not unduly affect the results in analytical studies.

It has long been recognised that estimate values for PUFA are difficult to harmonise when comparing nutritional databases. Frequent missing values of individual fatty acids and local conditions that determine variations in the chemical composition of fish and seafood (the primary source of PUFA) affect the comparability between food composition tables³¹. In addition, for the purposes of this research, we included updated Chilean data on the content of omega 3 and omega 6 fatty acids²⁰, which directly reflects the chemical composition of foods extracted in the country, so they may differ from those values reported when using the British table. These reasons may partly explain our findings of higher reported values in the American (Chilean) table and the lowest reported ICC among the macronutrients.

In our study, values for minerals showed the least agreement among micronutrients. Differences in vitamins and minerals have been confirmed in another Latin American study comparing the amount of minerals found

in vegetables and fruits to those values in the British and German food composition tables³². Food analyses from South America and Europe have demonstrated that the mineral content of foods is the most vulnerable to influence not only by the use of different methods, but also by environmental factors^{7,33}. Climatic conditions, light, temperature and soil characteristics can greatly affect the growth rates of plants and therefore the final content of minerals. This may explain why the agreement between the estimates of food composition in our study was lower for metals.

It should be noted that the Pearson correlation is a measure of association and not of agreement²⁴. In spite of this, it is the most common method used for assessing FFQ reproducibility. In fact, more than 90% of studies include only this analysis to compare dietary intake³, and it is wrongly inferred from a high correlation that 'the methods may be used interchangeably'. In our study, the ICC and Pearson correlation coefficient were not greatly different, but still the Pearson correlation would have provided a somewhat misleading reassurance in relation to iron and PUFA.

We conclude that there should be awareness that tables of food composition could be an important source of bias. Thus, comparison between studies that do not use the same table of food composition may be inappropriate. Methodological differences for determination of energy (including conversion factors), differences in supplementation policies as well as local environmental conditions influence nutritional values reported in tables of food composition. All of these characteristics can attenuate the level of association in aetiological studies. Taking on board all these caveats, the level of agreement between the two tables of food composition found in the present study was unexpectedly high.

Acknowledgements

We are indebted to Drs Patricia Bustos and Hugo Amigo for co-ordinating the data collection in Chile. We also thank Mrs Ana Cristina Pinheiro for assessing nutrient compositions based on the American (Chilean) table of food composition and the fieldworkers who administered the FFQ. This project was funded by a grant of the Wellcome Trust.

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